

1 Fecundity determines the extinction threshold in a Canadian assemblage 2 of longhorned beetles (Coleoptera: Cerambycidae)

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8 *Key words:* Cerambycidae, Emigration, Extinction threshold, Habitat loss, Reproductive rate, Spatial scale

9 Abstract

10 Reproductive rate has been suggested to have a positive effect on the amount of habitat loss a species can
 11 tolerate while emigration from habitat patches has been suggested to have both positive and negative
 12 effects. Forest fragmentation has been suggested to have negative effects on forest species. We determined
 13 the extinction threshold for 12 species of saproxylic (dead wood dependent) longhorned beetles (Coleop-
 14 tera: Cerambycidae) using trap catch data from Ontario, Canada. We also determined the maximum egg
 15 production of each species and whether they were likely to move outside of forest patches. We found a
 16 strong negative relationship between reproductive rate and the minimum habitat amount required for
 17 species presence. This relationship is obscured if the scale of investigation is not appropriate for the study
 18 organism. As well, species caught moving outside forest habitat had lower extinction thresholds than
 19 species not caught moving outside forest but this was not significant after accounting for reproductive rate.
 20 Fragmentation did not have an effect on the minimum habitat requirements. These relationships can inform
 21 predictions of which species will be most affected by habitat loss.

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25 Introduction

26 Habitat loss due to human activities is the main
 27 reason for the current loss of species (Terborgh
 28 1974; Ehrlich and Ehrlich 1981; Winchester 1997;
 29 Lande 1998). The conversion of natural areas to
 30 suit human activities such as urban development
 31 and agriculture (Leemans and Zuidema 1995)
 32 leaves less habitat for most other species (Walker
 33 1992; Turner et al. 1994). This reduction in habitat
 34 results in smaller and more isolated populations
 35 which are at greater risk of extinction due to sto-
 36 chastic demographic and environmental variation
 37 (Pimm et al. 1988; Raup 1991; Lande 1998).

Understanding how habitat loss affects the 38
 extinction risk of different species is therefore an 39
 important part of improving conservation efforts 40
 (Pimm et al. 1988; Pearson et al. 1999; With and 41
 King 1999; Fahrig 2001; Reed and Shine 2002). 42

An important aspect of a species' response to 43
 habitat loss is the possibility of a sudden increase 44
 in the probability of extinction at some critical 45
 amount of habitat (Lande 1987). This amount of 46
 habitat has been termed the extinction threshold 47
 (Lande 1987). This phenomenon has been found in 48
 spatially explicit models (Bascompte and Sole 49
 1996; With and King 1999; Fahrig 2001), and there 50
 is some empirical evidence for it as well (eg: Carl- 51

52 son 2000). Determining the critical amount of
 53 habitat that must be preserved to prevent different
 54 species from reaching the extinction threshold
 55 should be a key part of conservation research.
 56 Moreover, if species characteristics, such as life
 57 history traits or behaviour, correlate with the
 58 threshold, we could predict which species will be at
 59 risk at a given level of habitat loss without having
 60 to resort to detailed studies on each species. One
 61 problem in conservation is that in many cases the
 62 necessary remedial actions must be informed by
 63 detailed, species-specific studies (Doncaster et al.
 64 1996; Eriksson and Kiviniemi 1999). Easily
 65 obtained estimates of extinction risk could facili-
 66 tate conservation efforts (Davies et al. 2000;
 67 Duncan and Lockwood 2001).

68 Species characteristics that are likely to influence
 69 the risk of extinction include the intrinsic rate of
 70 population growth (Ehrensfield 1970; Bennet and
 71 Owens 1997; McKinney 1997) and the rate of
 72 movement between habitat patches (Terborgh
 73 1974; Lande et al. 1998; Huxel and Hastings 1999).
 74 Species with higher reproductive rates should re-
 75 bound from population declines more quickly
 76 (MacArthur and Wilson 1967). Small populations
 77 face a high risk of extinction (Pimm et al. 1988;
 78 Raup 1991; Boyce 1992; e.g. Hecnar and
 79 M'Closkey 1997) and increasing population num-
 80 bers quickly will reduce the time that the popula-
 81 tion is more vulnerable to subsequent stochastic
 82 events. Therefore species with greater reproductive
 83 rates should tolerate more habitat loss and so have
 84 a lower extinction threshold. A simulation study of
 85 factors affecting the extinction threshold predicted
 86 that reproductive rate had a very strong effect on
 87 the threshold amount of habitat (Fahrig 2001).
 88 Species with higher reproductive rate were pre-
 89 dicted to tolerate more habitat loss and so have a
 90 lower extinction threshold by this model. Vance
 91 et al. (2003) found evidence of this relationship in
 92 their study of forest breeding birds. They found a
 93 negative relationship between species reproductive
 94 rate and the amount of forest necessary for 50%
 95 probability of occurrence.

96 Emigration from habitat patches has also been
 97 predicted to affect the amount of habitat at the
 98 extinction threshold (Pagel and Payne 1996).
 99 Movement between habitat patches is needed for
 100 recolonization of patches in which local extinc-
 101 tions have occurred (Hanski et al. 1995; Britton
 102 et al. 2001). For example, Speight (1989) has

103 suggested that many forest remnants in the United
 104 Kingdom lack many species of saproxylic beetles
 105 even 200–300 years after being isolated because
 106 the distance to any source of colonizers is too great
 107 compared to the movement distances of the bee-
 108 tles. Dispersing individuals may also prevent local
 109 extinctions through a rescue effect (Brown and
 110 Kodric-Brown 1977). Therefore, the ability of a
 111 species to move through non-habitat areas is pre-
 112 dicted to restock or replenish declining popula-
 113 tions, thereby lowering the amount of habitat
 114 necessary for persistence.

115 In contrast, in a simulation study Fahrig (2001)
 116 found that increasing the emigration rate leads to
 117 an increase in the amount of habitat needed for
 118 population persistence. This increase in the
 119 extinction threshold was due to the increased
 120 mortality associated with movement through non-
 121 habitat (matrix) areas (Fahrig 2001). The amount
 122 of habitat at the extinction threshold was further
 123 increased by emigration when the probability of
 124 mortality in the matrix was increased. Therefore
 125 this model predicts that species with a high prob-
 126 ability of emigration from habitat patches will
 127 require more habitat in the landscape for persis-
 128 tence, or a higher extinction threshold, than spe-
 129 cies with a low probability of emigration. This
 130 difference will be greater in landscapes that contain
 131 a more hostile matrix. Because emigration has
 132 been predicted to have both negative and positive
 133 effects on population persistence, determining the
 134 effect size and direction for real species should be a
 135 conservation priority.

136 Habitat fragmentation within a landscape may
 137 also have an effect on the extinction threshold.
 138 Habitat fragmentation had a weak positive effect
 139 on the amount of habitat necessary for persistence
 140 in Fahrig's (2001) simulation study. Organisms
 141 that live in highly fragmented habitat may have a
 142 low probability of colonization and population
 143 rescue. In an extreme case, the patches in such a
 144 landscape may become completely isolated so that
 145 the individual patches are not colonized following
 146 local extinctions. Habitat fragmentation has
 147 probably been responsible for local extinctions of
 148 several species of ground beetles with low dispersal
 149 ability in heathland fragments (de Vries et al.
 150 1996).

151 Studies of the effects of habitat fragmentation
 152 are usually confounded by the effects of habitat
 153 loss (Fahrig 1997; Trzcinski et al. 1999). In most

154 studies, landscapes that are more fragmented also
 155 have less habitat. The isolation effects mentioned
 156 above are caused by the removal of the habitat
 157 between the remaining fragments, and not neces-
 158 sarily by fragmentation per se (Fahrig 1997). The
 159 correlation between habitat loss and habitat frag-
 160 mentation may lead to misguided management if
 161 fragmentation is perceived to be an important
 162 determinant of species persistence after habitat
 163 loss is accounted for when this is not in fact cor-
 164 rect. It is important to know whether the config-
 165 uration of habitat resulting from habitat loss can
 166 help mitigate that loss (Fahrig 1997). We therefore
 167 looked for the effects of fragmentation after
 168 accounting for habitat amount.

169 Determining whether extinction thresholds are a
 170 real pattern in species response to habitat loss is
 171 important if conservation efforts are to be suc-
 172 cessful because this would indicate that additional
 173 incremental losses of habitat may lead to extinc-
 174 tion rather than a small incremental reduction in a
 175 population with an associated small increase in the
 176 probability of extinction. As well, it is important
 177 that we understand how the characteristics of
 178 different species and landscapes affect the extinc-
 179 tion threshold in order to predict the effects of
 180 habitat loss and to guide conservation efforts. The
 181 purpose of this study was to test the effects of
 182 reproductive rate, emigration, and fragmentation
 183 on extinction thresholds using abundance data on
 184 saproxylic longhorned beetles (Coleoptera:
 185 Cerambycidae). We predicted that the effect sizes
 186 would be ordered reproductive rate > emigra-
 187 tion > fragmentation, following the simulation
 188 study results (Fahrig 2001). Because different
 189 species respond to habitat at different spatial scales
 190 (e.g.: Roland and Taylor 1997; Steffan-Dewenter
 191 et al. 2002; Holland et al. 2004) studies must be
 192 scaled appropriately for the study organisms
 193 (Addicott et al. 1987). We therefore looked for
 194 these relationships using the spatial scales of
 195 response of each species by measuring forest hab-
 196 itat at these scales.

197 **Methods**

198 The field work for this study was carried out in
 199 Ontario, Canada within 19 sampling areas located
 200 in the rural area to the west, south, and east of the
 201 city of Ottawa. Each 1 km² sampling area had 10

202 randomly located beetle trapping locations with a
 203 baited Lindgren funnel trap that ran for 4–
 204 5 months during the summers of 1999 and 2000.
 205 The data used in this study then is based on trap
 206 catches for approximately 1700 trap-months. De-
 207 tails of the sampling and the location of traps
 208 within the sampling areas can be found in Holland
 209 et al. (2004). We found the amount of habitat at
 210 the extinction threshold for 12 species of sapro-
 211 xylic (dead wood dependent) longhorned beetles
 212 using abundance data from Holland et al. (2004).
 213 The extinction threshold was defined as the lowest
 214 proportion of forest for which the species was
 215 present. The forest measurements were collected
 216 using digital topographical maps (National Capi-
 217 tal Commission 1999) within ArcView GIS (ESRI
 218 Corp., Redlands, California, USA).

219 The larvae of longhorned beetles mine galleries
 220 within either live or dead wood (Linsley 1954) and
 221 are confined to a single tree or piece of dead wood
 222 (Hanks 1999). The adult stage is brief in compar-
 223 ison to the larval stage, usually lasting only a few
 224 weeks (Safranyik and Moeck 1995). We only used
 225 species that have larvae that can develop within
 226 many different species of dead wood to ensure that
 227 the amount of forest was a good measure of the
 228 amount of habitat available to these species. Using
 229 only these species also avoided the possible con-
 230 founding effect of more specialized species being
 231 more prone to extinction (McKinney 1997; Kotze
 232 and O'Hara 2003).

233 A previous study demonstrated that each beetle
 234 species responds to forest habitat at a different
 235 characteristic scale (Holland et al. 2004). The
 236 characteristic scale of response refers to the scale
 237 (radius of circular area around a sampling site) at
 238 which the species responds most strongly to some
 239 aspect of its environment. In this study, it refers to
 240 the scale of forest measurement at which the
 241 relationship between the amount of forest habitat
 242 and the abundance of the species is strongest.
 243 Therefore, in this study the amount of forest that
 244 is available around each trapping site was mea-
 245 sured at the scale appropriate for that species, as
 246 determined in Holland et al. (in review).

247 The reproductive rate for each species was
 248 determined by dissecting 20 female beetles of each
 249 species and counting the number of developed
 250 eggs. Specimens from the Holland et al. (2004)
 251 study had been stored in 70% ethanol in the
 252 summers of 1999 and 2000. The eggs were in

253 excellent condition and easily counted when the
 254 specimens were dissected in autumn of 2002. We
 255 used the maximum number of eggs rather than the
 256 mean because some beetles would have already
 257 oviposited some eggs when caught and this was
 258 not possible to determine from the dissections.
 259 Reproductive output should also include genera-
 260 tion time (or number of clutches per year for some
 261 taxa). For the three species for which we were able
 262 to find this information in the literature the genera-
 263 tion time was 1 year. We assumed a generation
 264 time of 1 year for the remaining nine species and
 265 used the natural logarithm of the maximum egg
 266 production as a relative index (hereafter called
 267 reproductive rate) of population growth rate.

268 Cerambycidae species were classified as emi-
 269 grating outside forest habitat patches or remaining
 270 within forest habitat in an earlier study (Holland et
 271 al. in review). In this earlier study we used Lind-
 272 gren multiple funnel traps and flight intercept
 273 traps, both placed outside forest patches, to catch
 274 individuals emigrating from forest. This was done
 275 in areas where we were trapping within the forest
 276 patches as well and therefore knew which species
 277 were present. This allowed us to categorize 9 of the
 278 12 species used here as emigrating or not emi-
 279 grating. The other three species were not caught in
 280 the forest or matrix in areas with non-forest traps,
 281 and so these species could not be classified.

282 Each species was used as a data point in our
 283 analyses with extinction threshold as the response
 284 variable. To test the prediction of the effect of
 285 reproductive rate on extinction threshold we used
 286 linear regression analysis. To test the effect of
 287 emigration from forest on the extinction threshold
 288 we used a *t*-test (assuming unequal variance) to
 289 compare the proportion of forest at the extinction
 290 threshold for species caught moving outside forest
 291 habitat and those caught only within forest. We
 292 tested the effects of these two variables separately
 293 because the lack of movement data for three of the
 294 species meant these could not be included in a
 295 multiple linear regression. Using a multiple linear
 296 regression with the remaining nine species would
 297 have resulted in very low power to detect an effect
 298 of reproductive rate, but we did use multiple
 299 regression analysis to check the emigration results
 300 after taking reproductive rate into account.

301 A possible confounding factor in the analysis
 302 of the effect of emigration is the difference in
 303 commonness among the different species. If all

species are caught moving outside forest in direct
 proportion to their relative numbers and not
 because there is a real difference in movement,
 this would lead to the more common species
 being assigned to the “species moving” category
 despite the lack of a difference. To see if this was
 occurring we conducted a *t*-test (assuming
 unequal variance) to compare the mean total
 number of individuals caught in the moving and
 not moving categories.

We used the Effective Number of Habitat Pat-
 ches (ENHP) to measure forest fragmentation
 (Jaeger et al. in review). The formula for this is
 $ENHP = 1/(\sum(A_i/A_t)^2)$, where A_i is the area of
 habitat patch i , and A_t is the total area of the study
 region. In our case, A_t is the area encompassed by
 a circle with radius equal to the characteristic scale
 of response of the species to forest. The ENHP is
 the reciprocal of the degree of coherence (C);
 conceptually, C measures the probability that any
 two points randomly placed in the region will
 occur in the same habitat patch (Jaeger 2000). We
 wanted to include small treed patches and fence-
 rows when measuring fragmentation, so digital
 1:15000 colour air photographs (City of Ottawa
 2000) were used to digitize all treed patches that
 were not included in the original digital topo-
 graphical maps (National Capital Commission
 1999). M. Burrell digitized these features using
 ArcView. Fencerows were digitized as continuous
 features whenever the canopy gaps were less wide
 than twice the canopy width at the gap. All treed
 fencerows and small treed patches within 2 km of
 the trapping sites were digitized. We then created
 two separate forest habitat themes with which to
 measure fragmentation: one with all originally
 mapped forest patches plus the smaller patches
 that we digitized, and one that further included all
 treed fencerows. Within each theme we merged all
 contiguous and overlapping patches and fence-
 rows. Therefore two patches joined by a fencerow
 were considered a single patch. Patches that
 appeared separate within the 2 km radius, but that
 were actually joined by some connection beyond
 the 2 km line, were treated as a single patch. The
 ENHP was then calculated by measuring the
 proportion of the region that each forest patch
 represented within a given radius of each of the
 trapping sites using a custom ArcView script. The
 fragmentation with and without fencerows was
 calculated at spatial scales of 20–200 m in 20 m

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355 increments, and 400 m to 2 km in 200 m incre-
 356 ments. We compared the range of values of the
 357 fragmentation values across all scales both with
 358 and without fencerows included. Both these mea-
 359 sures of fragmentation had a relatively large range
 360 of values at 1400 m so we used this as the scale at
 361 which to compare the effects of habitat amount
 362 and fragmentation.

363 To test for effects of habitat amount and
 364 fragmentation we used multiple logistic regression
 365 with occurrence as the response variable and
 366 habitat amount and fragmentation at 1400 m as
 367 predictor variables. We included species as a class
 368 variable. The measures of fragmentation with and
 369 without fencerows were highly correlated (Pear-
 370 son $r=0.8657$) so we used 2 models, each
 371 including one of these. These tests had inflated
 372 power because they used data from all 190 plots
 373 despite the fact that the measured predictor
 374 variables are not from spatially independent areas
 375 at the 1400 m scale. We adjusted the results for
 376 this by applying a correction factor to the stan-
 377 dard error of the estimates before calculating the
 378 Wald statistic and testing for significance. At
 379 1400 m, randomly sampling spatially independent
 380 sites (separated by at least 2800 m) yields an
 381 average of 18 sites possible (Figure 1). We
 382 therefore multiplied the standard error values by
 383 $(190/18)^{1/2}=3.249$. We also checked the signifi-
 384 cance of habitat amount-fragmentation interac-
 385 tion terms.

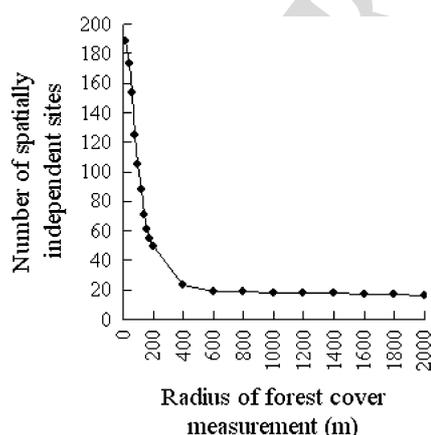


Figure 1. The relationship between the spatial scale at which we measured habitat and the number of sites that remained spatially independent, or did not have overlapping areas of forest cover measurement.

Results

Figure 2 shows the abundance of each species across the range of forest habitat proportions, with this proportion measured at the characteristic scale of response to forest habitat for each species (Holland et al. 2004). The proportion of forest at the extinction thresholds for the 12 species ranged from 0.0476 to 0.9897 (Table 1) forest cover. The maximum number of eggs per female within a species ranged from 9 to 236, corresponding to reproductive rates of 2.20 to 5.46 (Table 1). Larger beetle species tended to have greater egg counts. To test this we performed a linear regression between mean body length data from Holland et al. (in review) and the maximum egg count with the latter as the response variable. There was a significant positive relationship between maximum egg count and body size ($R^2=0.555$, $F=12.5$, $df=11$, $p=0.0054$). There was a significant negative effect of reproductive rate on the extinction threshold (Figure 3a, $R^2=0.617$, $F=16.1$, $df=11$, $p=0.0025$). Species with higher reproductive output were able to persist in areas with less forest habitat, as predicted. This relationship explained 61.7% of the variance in the extinction threshold values.

The mean number caught and associated standard errors of the emigrating and not emigrating categories were 40.0 ± 11.0 and 46.1 ± 13.8 respectively. This test clearly showed that the species assigned to the moving category were not more common than those we didn't catch moving ($t=0.371$, $df=5$, $p>0.3$), so that there is a real difference in the probability of emigration from forest between the two groups. The t -test revealed a significant difference ($t=3.935$, $df=7$, $p<0.01$) in the extinction thresholds between the species we caught moving and those not caught moving outside forest habitat (Figure 4). The species that were caught moving had lower extinction thresholds. To ensure that this result held after accounting for reproductive rate we used a multiple linear regression with emigration and reproductive rate as predictors of the extinction threshold. Despite the lowered power of this analysis reproductive rate remained a strong predictor of the extinction threshold amount of habitat, but emigration was no longer significant.

The logistic regression with fragmentation measured including fencerows was not significant

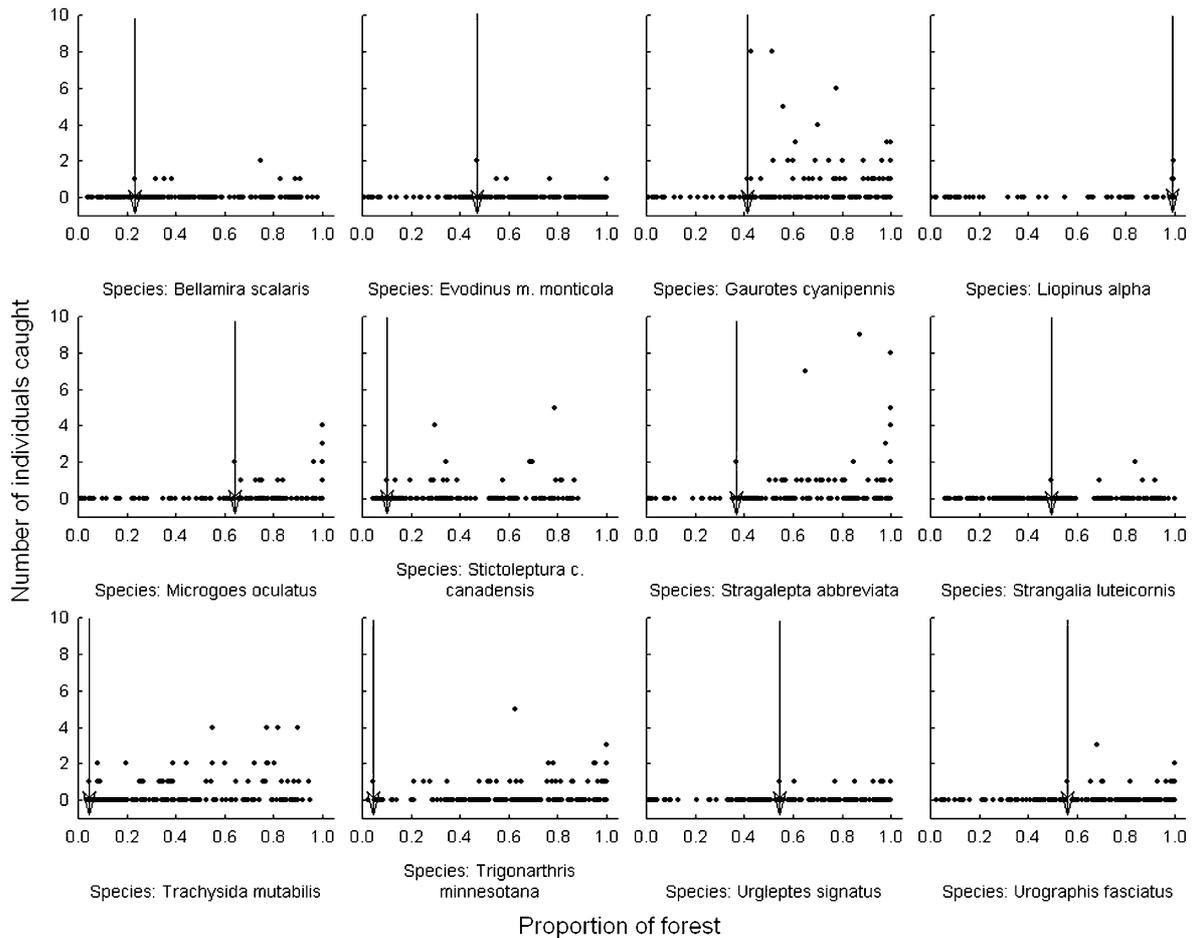


Figure 2. Determining the extinction threshold for the Cerambycidae species. The number of individuals caught at our sampling sites was plotted against the proportion of forest cover around the sites. The forest cover was measured at the characteristic scale of response for each species (Table 1). The arrows indicate the extinction threshold amount of habitat for each species.

Table 1. Values of the predictor variables and the extinction thresholds.

Species	Max. eggs	Repr. rate	Emigrate?	Extinction threshold	Scale of response (m)
<i>Bellamira scalaris</i> (Say)	236	5.46	N	0.231	1000
<i>Evodinus m. monticola</i> (Rand.)	37	3.61		0.471	160
<i>Gaurotes cyanipennis</i> (Say)	101	4.62	N	0.413	160
<i>Liopinus alpha</i> (Say)	10	2.30		0.990	20
<i>Microgoes oculatus</i> (LeC.)	20	3.00	N	0.641	60
<i>Stictoleptura c. canadensis</i> (Oliv.)	160	5.08	Y	0.101	1600
<i>Stragalepta abbreviata</i> (Germ.)	40	3.69	N	0.369	120
<i>Strangalia luteicornis</i> (F.)	9	2.20		0.497	800
<i>Trachysida mutabilis</i> (Newm.)	80	4.38	N	0.048	1200
<i>Trigonarthris minnesotana</i> (Csy.)	152	5.02	Y	0.048	200
<i>Urgleptes signatus</i> (LeC.)	14	2.64	N	0.546	140
<i>Urographis fasciatus</i> (DeG.)	58	4.06	N	0.559	180

The extinction thresholds were determined as per Figure 2, with the amount of forest cover measured at the characteristic scale of response of each species.

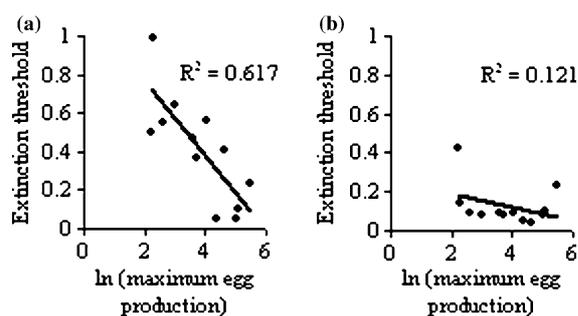


Figure 3. Relationship between reproductive rate and the extinction threshold. In a, each of the twelve species had the extinction threshold determined at the spatial scale that it responds most strongly to forest habitat, or the characteristic scale of response to forest habitat. The equation of the best fit line was $\text{ext.thresh.} = -0.1929(\ln(\text{max. egg. prod.})) + 1.1499$. In b, each of the 12 species had the extinction threshold determined by plotting the abundance against the proportion of forest within 1 km of the sampling site for each species. This relationship is no longer significant when done at this scale.

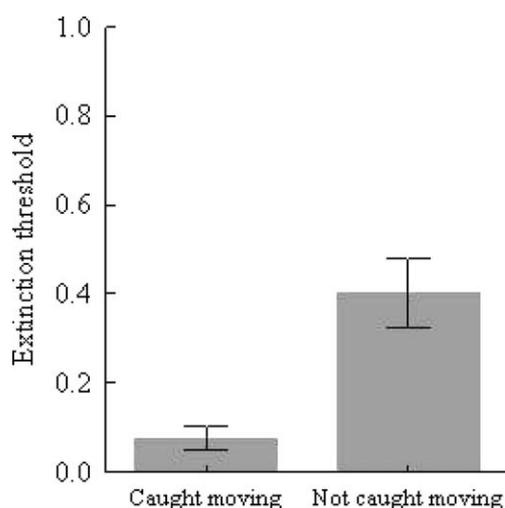


Figure 4. Comparison of extinction thresholds of species caught outside forest habitat ($N=2$) and species not caught outside of forest habitat ($N=7$). The species caught outside forest habitat had a lower mean habitat amount at the extinction threshold. Emigration was not significant when included in a multiple linear regression with reproductive rate as a second predictor variable.

436 even before correcting for the number of spatially
 437 independent trapping locations so we assumed that
 438 the fragmentation measured without these features
 439 was a better predictor of occurrence. The habitat
 440 amount-fragmentation interaction terms were not
 441 significant in either model. After adjusting the

Wald statistics for the possible number of spatially
 independent sites, fragmentation measured with-
 out fencerows was not significant (Wald = 1.0566,
 $df = 1, p > 0.3$) while habitat amount was still a
 marginally significant positive predictor of occur-
 rence (Wald = 3.8075, $df = 1, p = 0.051$).

Discussion

The results support the prediction that reproductive
 rate has a strong negative effect on the extinction
 threshold amount of habitat. Species with low
 reproductive rates were only found in landscapes
 with a large proportion of forest remaining,
 whereas species with higher reproductive rates were
 found in landscapes with less forest habitat
 remaining. This result implies that, as habitat is lost
 in a landscape, species with lower reproductive rates
 experience local extinctions before species with
 higher reproductive rates. This agrees with the
 empirical study of forest breeding birds by Vance
 et al. (2003). Similar to our study, their results imply
 that species with lower reproductive rates will be-
 come locally extinct first as habitat is lost.

Using an arbitrary scale to determine the
 extinction threshold may have caused us to miss
 the reproductive rate-extinction threshold rela-
 tionship. To test this idea we used an arbitrary
 1 km scale (radius around each trapping site) to
 measure forest and plotted the abundance of each
 species against the proportion of forest within a
 1 km radius. We again considered the extinction
 threshold to be the minimum proportion of forest
 at which a species occurred. A regression of the
 extinction thresholds against the reproductive rate
 of the species reveals a much weaker trend
 (Figure 3b) that is not statistically significant
 ($R^2 = 0.121, F = 1.378, df = 11, p > 0.2$). The vari-
 ance explained by this regression is about one-fifth
 that explained by the relationship that we found
 by using the characteristic scale of response for
 each species. The non-significance of the results
 obtained using the same arbitrary 1 km scale for
 all species would have led to the conclusion that
 reproductive rate does not have an effect on the
 extinction threshold. The difference between this
 result and that found using the characteristic scale
 of response for each species underscores the
 importance of conducting studies at an appropri-
 ate spatial scale (compare Figure 3a and b).

490 Vance et al. (2003) used the same spatial scale
 491 for all species in their study of forest breeding
 492 birds. Interestingly, the variance explained by the
 493 reproductive rate-habitat necessary regression in
 494 their study ($R^2 = 16$) is very close to the value we
 495 obtained when using a constant scale for all species
 496 ($R^2 = 0.121$). This suggests that the magnitude of
 497 the reproductive rate effect on the extinction
 498 threshold amount of habitat may be similar in
 499 different taxa. It seems likely that they would have
 500 found an even stronger effect if the habitat neces-
 501 sary for 50% probability of persistence had been
 502 measured at the characteristic scale of response for
 503 each species. Although the variance explained in
 504 their study is only slightly higher than in our re-
 505 sults with a constant scale, they did find a signifi-
 506 cant effect of reproductive rate this way, probably
 507 because they had a larger number of species
 508 ($n = 41$).

509 Larger beetles tended to have a greater maxi-
 510 mum egg count. This raises the possibility that
 511 the mean size of the species is largely responsible
 512 for determining the extinction threshold rather
 513 than fecundity. We did not have any *a priori*
 514 reason to suspect that body size would influence
 515 the extinction threshold and so did not include
 516 this as a predictor. In fact, studies that do use
 517 body size as a predictor of proneness to extinc-
 518 tion tend to use this as a composite indicator of
 519 other variables such as fecundity. In our study,
 520 the size-reproductive rate is positive, opposite to
 521 most studies of other taxa. We conclude that
 522 larger beetles tended to have greater reproductive
 523 and therefore lower extinction thresholds; the
 524 threshold amount of habitat is determined by
 525 reproductive rate, not body size.

526 Our results indicate that species prone to emi-
 527 gration out of forest habitat may have a lower
 528 extinction threshold than species that do not move
 529 out of forest habitat. However this result did not
 530 hold in the (admittedly low power) multiple
 531 regression with reproductive rate. It is possible
 532 that the increase in non-forest area does not lead
 533 to an increase in mortality during movement for
 534 the species that move out of forest. The idea that
 535 non-habitat matrix is completely hostile is an
 536 ideological artifact of island biogeography (D'Eon
 537 2002) and does not always apply to species moving
 538 between terrestrial habitat patches (Huxel and
 539 Hastings 1999). As adults, many species of long-
 540 horned beetles eat pollen and nectar (Yanega

1996) and are commonly found on flowers of
 shrubs and herbaceous vegetation in fencerows
 (Samways 1994) and other non-forest habitats.
 For species that are commonly found feeding as
 adults within non-forest habitat, movement into
 non-forest areas may actually increase survival of
 these species. Some species have evolved either
 mimicry or cryptic colouration (Yanega 1996),
 suggesting that they are at least partially adapted
 to predators. Laurance (1991) found that the
 extinction proneness of 66 species of rainforest
 mammals was negatively related to their tolerance
 of the matrix.

It is also likely that the species that move outside
 forest to obtain floral resources are more likely to
 recolonize forest patches that experience local
 extinctions than species that do not readily move
 through non-forest areas. Females of the species
 that move outside forest must locate suitable larval
 habitat for laying eggs. Shibata (1987) has found
 evidence to suggest that species of longhorned
 beetles that must move to find food and then
 oviposition sites disperse farther than species that
 are not required to do so. It is very likely that in
 these species, females lay eggs in a forest patch
 other than the one they are from. Therefore, the
 species that leave forest habitat likely have higher
 rates of patch recolonization and higher rates of
 rescue from low numbers than species that do not
 move outside forest habitat. Unfortunately, we
 were not able to reliably separate the effects of
 reproductive rate and emigration to test the effect
 of emigration.

Our coarse classification of movement propen-
 sity may mask more complex trends in the effect of
 dispersal on the extinction threshold. Thomas
 (2000) used data on butterflies grouped into three
 movement classes and found that species of inter-
 mediate mobility have shown the greatest recent
 declines, probably because these species tended to
 move out of habitat but not locate another suit-
 able habitat location. The movement classes that
 Thomas (2000) made use of were based on actual
 movement distances. There is no information on
 movement for most species of cerambycids, and
 our flight intercept trapping in non-forest was an
 attempt to gain some of this knowledge.
 Improvements in dispersal monitoring techniques
 may yet allow for the collection of detailed
 movement data on cerambycids. There are at least
 two ways that we could reconcile our classification

592 with the results of Thomas (2000). One possibility
 593 is that our binary emigrate/don't emigrate classi-
 594 fication is appropriate and that the species we
 595 caught moving are comparable to lumping Tho-
 596 mas's long and medium range dispersers. If this
 597 scheme is appropriate then our results may not
 598 coincide with those for the butterflies. A second
 599 possibility is that there are some species that move
 600 so little that they were present but not caught even
 601 by our traps located within forest habitat. We
 602 could then classify species into three classes: spe-
 603 cies caught moving outside forest (long-distance
 604 dispersers), species caught only within forest
 605 (medium-range dispersers), and species present but
 606 not caught at all (non-dispersers). The non-dis-
 607 persers may be unaffected by processes outside a
 608 very small area around their host log and so not
 609 face extirpation from habitat loss. If it is accurate
 610 then it is possible that our species are reacting to
 611 habitat loss in a manner similar to Thomas's
 612 butterflies if the unobserved non-dispersers have a
 613 low extinction threshold. Of course this idea is
 614 impossible to test with our current data since it
 615 relies on an unobserved class of species.

616 Our results show that fragmentation is not
 617 important in determining the occurrence of these
 618 species. In this test we gave considerable weight to
 619 the fragmentation variable by conducting the
 620 multiple logistic regression at a spatial scale where
 621 the range of fragmentation values was greatest. In
 622 fact they had a much greater range than the forest
 623 habitat amount values at this scale, but habitat
 624 amount was still a marginally significant predictor
 625 of occurrence while fragmentation was not. This
 626 result agrees with other studies that have found
 627 that habitat amount is much more important than
 628 habitat fragmentation (McGarigal and McComb
 629 1995; Trzcinski et al. 1999). Conservation efforts
 630 directed at these species should focus on forest
 631 habitat retention, and should not be misguided by
 632 the idea that configuration can ameliorate habitat
 633 loss (Fahrig 1997).

634 If the distance that edge effects extend into the
 635 forest is large relative to the characteristic scale of
 636 response, taxa that respond to habitat at relatively
 637 small scales may seem to have very high extinction
 638 thresholds only because they are 'interior species'.
 639 This would cause the species to only occur in areas
 640 that are surrounded by close to 100% forest
 641 'habitat', but only because the edge is not properly
 642 included in the measurement of the amount of

643 habitat. Researchers studying interior species
 644 should be aware of the possibility of such false
 645 extinction thresholds. While such responses are
 646 more correctly viewed as a response to habitat
 647 type, they could still provide information on suit-
 648 able conditions for the species of concern. This
 649 was probably not a problem in the present study,
 650 as none of the longhorned beetles in this study
 651 have been described as interior species to our
 652 knowledge. Eight species considered here
 653 responded at small scales of up to 200 m, but
 654 within these species the observed extinction
 655 thresholds ranged from 0.0478 to 0.9897. The only
 656 species with a small characteristic scale of response
 657 and a high extinction threshold was *Liopinus al-*
 658 *pha*.

659 We know that the location and size of forest
 660 patches in our 29 study areas has been relatively
 661 constant during the past 25–30 years (Contreras
 662 2002). However, much of the forest cover in the
 663 study region around Ottawa, Ontario was
 664 removed during intensive logging during the early
 665 and mid-1800's (Keddy 1993). It is possible that
 666 the intervening period has not been long enough
 667 for the cerambycid species in the area to come to
 668 equilibrium with the resulting habitat composi-
 669 tion. If there is an extinction debt (sensu Tilman
 670 et al. 1994) this could affect our determinations of
 671 the extinction threshold amount of habitat. The
 672 affected species could be extant but lost from some
 673 areas in the future even without further addition
 674 loss of forest habitat. In such a scenario our esti-
 675 mates of the extinction threshold amount of habi-
 676 tat would be low; species would actually require
 677 more habitat than is suggested. Further, we do not
 678 know if the extinction thresholds should all be
 679 shifted by the same amount for the different spe-
 680 cies. Using insects can help to minimize the pos-
 681 sibility of a time lag in response to habitat loss if
 682 species with short generation times are used. Given
 683 additional data to make comparisons, we could
 684 check for the possibility of an extinction debt by
 685 looking for a shift within the assemblages towards
 686 more rare species (Hanski and Ovaskainen 2002).

687 Our findings show that the minimum amount of
 688 habitat required for occurrence depends on life
 689 history attributes of the species. Species with
 690 greater reproductive rate were able to persist in
 691 areas containing less forest habitat than species
 692 with lower reproductive rates. Relationships
 693 between species and their environment are ob-

694 scured if the scale of investigation is not relevant to
 695 the species being studied. We suggest that the first
 696 step in such studies should be careful consideration
 697 of the research questions and the determination of
 698 the spatial scale at which the studied relationship is
 699 relevant.

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